Women’s preferences, willingness-to-pay, and predicted uptake for single-nucleotide polymorphism gene testing to guide personalized breast cancer screening strategies: a discrete choice experiment

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Background: Single-nucleotide polymorphism (SNP) gene test is a potential tool for improving the accuracy of breast cancer risk prediction. We seek to measure women’s preferences and marginal willingness-to-pay (mWTP) for this new technology.

Materials and methods: We administered a discrete choice experiment (DCE) to English-speaking Singaporean women aged 40–69 years without any history of breast cancer, enrolled via door-to-door recruitment with quota sampling by age and ethnicity. DCE attributes comprise: 1) sample type (buccal swab and dried blood spot), 2) person conducting pretest discussion (specialist doctor, non-specialist doctor, and nurse educator), 3) test location (private family clinic, public primary-care clinic, and hospital), and 4) out-of-pocket cost (S$50, S$175, and S$300). Mixed logit model was used to estimate the effect of attribute levels on women’s preferences and mWTP. Interactions between significant attributes and respondent characteristics were investigated. Predicted uptake rates for various gene testing scenarios were studied.

Results: A total of 300 women aged 52.6±7.6 years completed the survey (100 Chinese, Malay, and Indian women, respectively). Sample type (P=0.046), person conducting pretest discussion, and out-of-pocket cost (P<0.001) are significantly associated with going for SNP gene testing. Women with higher income and education levels are more willing to pay higher prices for the test. Preferences in terms of mWTP across ethnic groups appear similar, but Chinese women have greater preference heterogeneity for the attributes. Predicted uptake for a feasible scenario consisting of buccal swab, pretest discussion with nurse educator at the hospital costing S$50 is 60.5%. Only 3.3% of women always opted out of the SNP gene test in real life. Reasons include high cost, poor awareness, and indifference toward test results.

Conclusion: SNP gene testing may be tailored according to individual preferences to encourage uptake. Future research should focus on outcomes and cost-effectiveness of personalized breast cancer screening using SNP gene testing.

Keywords: single-nucleotide polymorphisms, gene testing, personalized breast cancer screening, precision medicine, women’s preferences, willingness-to-pay, predicted uptake, discrete choice experiment

Introduction
Breast cancer is a leading cause of morbidity and mortality in women worldwide. In 2015, it reached 2.4 million cases in incidence while accounting for 15.1 million
disability-adjusted life years.\textsuperscript{1} Mammography is currently the best screening tool for breast cancer and is linked to reduction in breast cancer mortality.\textsuperscript{2} However, with the current one-size-fits-all age-based screening, false positives and overdiagnosis are common.\textsuperscript{3} Hence, an optimized screening strategy based on individual woman’s risk estimates has been advocated. It is suggested as being more efficient and effective than the age-based approach.\textsuperscript{4} Under this strategy, women at higher risk may undergo screening more frequently, while women at lower risk may be screened less frequently, allowing for earlier detection of breast cancer in the former and reducing harms associated with excessive screening in the latter.

Several risk stratification tools such as the Tyrer-Cuzick Model\textsuperscript{5} and the Breast Cancer Screening Consortium Risk Calculator\textsuperscript{6} are available. Researchers have recommended these tools to be incorporated in screening decisions.\textsuperscript{7} However, they are limited by moderate discriminatory accuracy.\textsuperscript{8,9} Recent genome-wide association studies have discovered multiple single-nucleotide polymorphisms (SNPs) in the population that are associated with breast cancer.\textsuperscript{10–15} Although each SNP confers only low to moderate risk, they occur at higher frequencies than higher-penetrance BRCA mutations.\textsuperscript{16} When combined into a polygenic risk score (PRS), these SNPs improved the discriminatory accuracy of the existing risk assessment models\textsuperscript{17,18} and refined risk classification by shifting women at borderline low and high risk into low- and high-risk groups, respectively.\textsuperscript{19} Hence, risk stratification using PRS information may potentially inform risk-based screening strategies.\textsuperscript{20}

Furthermore, little is known about women perspectives regarding SNP gene testing for breast cancer risk.\textsuperscript{21–23} Increasingly, regulators such as the United States Food and Drug Administration as well as the European Medicines Agency have been working on incorporating risk–benefit perspectives of patients in regulatory assessments.\textsuperscript{24,25} Likewise, in the health technology assessment process, the engagement of end-users broadens the perspective of recommendations given to decision makers\textsuperscript{26} and provides a real-world understanding of the benefits and risks of health technologies and possible future uptake.\textsuperscript{27}

Discrete choice experiments (DCEs) have been used to assess personal preferences, risk–benefit trade-offs and marginal rates of substitution (eg, marginal willingness-to-pay [mWTP]) in breast cancer screening.\textsuperscript{28–31} This led to better understanding of women’s trade-offs between benefits and costs of screening including intangible ones like false positives and overdiagnosis,\textsuperscript{28,31} importance of process attributes such as screening time and desire for privacy in influencing screening uptake\textsuperscript{32} and heterogeneity in preferences according to their socioeconomic background and health status.\textsuperscript{33} DCEs conducted in cancer gene testing\textsuperscript{32–35} found that respondents value risk information and surveillance advice much more than assistance with decision making,\textsuperscript{12} prefer combinations of test characteristics that reflect future genomic testing more closely than current genetic testing\textsuperscript{33} and value cost and privacy more than reduction in false negatives.\textsuperscript{35} Yet, preferences for a one-time SNP gene testing to guide personalized screening strategies in breast cancer have not been extensively investigated. Hence, we aimed to estimate women’s preferences and mWTP for the various aspects of SNP gene testing and to predict uptake in a multi-ethnic Singapore population using DCE.

**Material and methods**

**Recruitment and sample size**

The study was conducted among women in the Singapore general population in December 2016. Multistage cluster sampling was employed (accompanying details in Supplementary materials). Screening-age women between 40 and 69 years were selected based on pre-specified ethnic quotas. Other eligibility criteria include Singapore citizenship and no prior history of breast cancer. Face-to-face surveys were conducted, and responses were recorded on tablet. All respondents provided written informed consent prior to participation. The study was approved by the National Healthcare Group Domain Specific Review Board (reference number: 2016/00184). Based on the Orme’s Rule of Thumb,\textsuperscript{36} the minimum sample size required is 75. We aimed for 100 women in each of the three ethnicities, with a total of 300 women. This was planned before the publication of a practical guide on sample size requirements by de Bekker-Grob et al.\textsuperscript{37}

**Ethics approval**

All procedures performed in studies involving human participants were in accordance with the ethical standards of the institutional and/or national research committee and with the 1964 Helsinki Declaration and its later amendments or comparable ethical standards.

**Study measures**

Respondents were taken through a two-page explanation on SNP gene testing before completing the DCE. Sociodemographic characteristics, family history of breast cancer, and mammogram screening history were solicited. Respondents received S$30 (US$23) shopping vouchers as compensation for their time and effort.
Women’s preferences and willingness-to-pay for SNP gene test

DCE construction

As cost was included as one of the characteristics, mWTP can be calculated, to resemble real-world situations where money is spent to obtain improvement in certain characteristics of a good or service. Our DCE was developed based on good research practices. Details can be found in Supplementary materials. Input from screening-age women in a qualitative study, clinicians, health service researchers, and the literature informed the choice of attributes and their associated levels for the DCE (Table 1). Notably, women in the study expected at least 90% test accuracy, defined as “how accurate the SNPs gene test is, in assessing risk of developing breast cancer,” before they would consider doing the test. They were less concerned about overdiagnosis or false positives. Due to the lack of realistic estimates of risk stratification benefits, for example, proportion of late-stage cancers diagnosed and biopsy rate, we did not include them in the DCE.

Respondents were presented with a series of choice tasks and asked to choose one out of two alternatives labeled A and B in each task. They were then asked if they would really go for it in real life. This is known as a dual-response none question (closed-ended question in Part B of Figure 1, sample choice task). It prevents overestimation of screening uptake as preference for one option over another does not mean that the preferred option will be endorsed. Reasons for not going for the test were solicited in the open-ended question in Part B.

Sawtooth Lighthouse Studio 6.4.4. (Sawtooth Software, Orem, UT, USA) was used to generate 100 unique questionnaire versions, replicated in each of the three ethnicities, each containing one fixed task and 10 choice tasks randomly generated using balanced overlap method. It allowed for moderate attribute level overlap across alternative scenarios and better discrimination when respondents use non-compensatory rules in making decisions between choice alternatives. Although this reduced statistical efficiency as the number of levels that were directly compared was reduced, it lightened the cognitive burden on the respondents. A subset of the full-choice design (ie, fractional factorial design) was sampled for each respondent, while ensuring level balance

### Table 1 Attributes and levels of SNPs gene testing

<table>
<thead>
<tr>
<th>Attribute</th>
<th>Levels</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample type</td>
<td>1 Buccal swab 2 Dried blood spot</td>
</tr>
<tr>
<td>Person conducting pretest discussion</td>
<td>1 Specialist doctor 2 Nurse educator trained in genetic counseling 3 Non-specialist doctor</td>
</tr>
<tr>
<td>Test location</td>
<td>1 Private family clinic 2 Hospital 3 Public primary-care clinic</td>
</tr>
<tr>
<td>Out-of-pocket cost</td>
<td>1 S$50 (US$38) 2 S$175 (US$134) 3 S$300 (US$229)</td>
</tr>
</tbody>
</table>

Note: US$1 is approximately S$1.31.

Abbreviations: SNP, single-nucleotide polymorphism; S$, Singapore dollars.

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**Figure 1** Example of a random choice task.

Abbreviations: SNP, single-nucleotide polymorphism; S$, Singapore dollars.
Women are willing to pay in exchange for a more preferred
the groups. mWTP of the overall study population was also
cient and allowing direct comparison of preferences across
normalization of coefficients by eliminating the scale coef
marginal rates of substitution between out-of-pocket cost and
Preferences among the ethnic groups were compared using
respondent characteristics (Table 2). The final model was built using forward
selection, where all attributes (main effects) were included,
only interaction terms between significant attributes and

Statistical analysis
Women's preferences
Respondent characteristics were described using counts and
percentages for categorical variables and mean (SD) and
median (interquartile range) for continuous variables. DCE
data were estimated using mixed logit model in the prefer-
ence space to calculate the part-worth utilities (PWUs) or
preference weights for all attribute levels using Stata 14.2
(StataCorp LLC, College Station, TX, USA). Mixed logit was
chosen over other models, for example, conditional logit, as
it accounts for preference heterogeneity around parameter
estimates among respondents, which was observed in our
data as indicated by a likelihood ratio test for the joint sig-
ificance of the standard deviations. A main effects model
was estimated, with the attribute levels for “sample type,”
“person conducting pretest discussion,” and “test location”
being effects-coded and “out-of-pocket cost” being coded as
a continuous variable, as our data showed a linear relation-
ship with PWU when “out-of-pocket cost” was coded as
categorical. Alternative-specific constants (ASC) for choos-
ing the left-sided alternative and not to go for the test in real
life (closed-ended question in Part B) were also included
in the main effects model. Interactions between significant
attribute levels and respondent characteristics were studied
to determine the latter’s effect on preferences for the attribute
levels. The first category of each characteristic is the refer-
ence level (Table 2). The final model was built using forward
selection, where all attributes (main effects) were included,
and only interaction terms between significant attributes and
respondent characteristics (P<0.05) were considered.

mWTP
Preferences among the ethnic groups were compared using
marginal rates of substitution between out-of-pocket cost and
other attributes, which is the same as mWTP. This resulted in
normalization of coefficients by eliminating the scale coef-
ficient and allowing direct comparison of preferences across
the groups. mWTP of the overall study population was also
studied, as it indicates the additional out-of-pocket cost that
women are willing to pay in exchange for a more preferred
level within an attribute. These were derived by estimating
the mixed logit model in WTP space rather than preference
space, with the cost coefficient being lognormally dis-
tributed and all other attribute coefficients being normally
distributed. To assess mWTP heterogeneity, individual-level
parameters were obtained post-estimation to construct prob-
ability density and cumulative density functions.

Predicted uptake
In budget impact analysis and program planning, one will be
interested in predicted uptake. Details on the calculation of
predicted uptake using PWU is in Supplementary materials.
The mean indirect utility over all possible testing alternatives
is fixed at zero. This means that the predicted uptake rates
are estimated relative to the uptake rate of the grand mean.
Alternatives include a realistic base case scenario reflecting
current genetic testing (buccal swab, specialist doctor, hos-
pital, and S$175) and other options that are feasible, cheaper
or pricier, and most preferred by the study population.

Dual-response none
Respondent choice in the closed-ended question of Part B
(yes or no) was analyzed together with Part A, forming the
DCE data. Proportion of women who chose not to undergo
SNP gene test in real life at least once or all the time were
described using counts and percentages. Responses to the
open-ended question of Part B, that is, reasons for not going
for the test were analyzed for commonality and coded accord-
ingly, to derive categories of common responses.

Results
Respondent characteristics
A total of 300 women completed the survey (300 out of
598 eligible dwelling units; 50.2% response rate). On aver-
age, the respondents took 22.6 minutes (SD: 10.5 minutes)
to complete it. Reasons for non-participation included
refusal and non-availability for re-appointment. The aver-
age age of respondents was 52.6 years (SD 7.6 years). Other
respondent characteristics are presented in Table 2.

Women’s preferences and mWTP (n=300)
Parameter estimates for the attribute levels are shown in
Table 3. All levels of sample type (P=0.046), person con-
ducting pretest discussion, and out-of-pocket cost are statistically
significant (P<0.001), while all levels of test location are
nonsignificant. Respondents preferred an SNP gene test that
is less invasive (buccal swab), involving a pretest discussion
with specialist doctor at a private family clinic. The least
Women's preferences and willingness-to-pay for SNP gene test

preferred combination is a more invasive test (dried blood spot) involving a non-specialist doctor at a public primary-care clinic. There is a positive association between choosing the left-sided alternative and opting for SNP gene testing ($P < 0.001$), reflecting a left-sided bias. However, a comparison between main effects models including and excluding the ASC “Choose left” found that the preference ranking within each attribute is similar except for the test location attribute (Table S1), which is not significantly associated with going for SNP gene testing. In the model that included ASC “Choose left,” private family clinic is most preferred, followed by hospital and public primary-care clinic. In the model that excluded ASC “Choose left,” private family clinic is most preferred, followed by public primary-care clinic and hospital.

Interactions between respondent characteristics and attribute levels are also shown in Table 3. Among four attributes and two ASCs, the interactions are significant

### Table 2 Self-reported respondent characteristics (n=300)

<table>
<thead>
<tr>
<th>Sociodemographic characteristics</th>
<th>Total number</th>
<th>Chinese</th>
<th>Malay</th>
<th>Indian</th>
<th>P-value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Total number of respondents</td>
<td>300 (100.0)</td>
<td>100 (33.3)</td>
<td>100 (33.3)</td>
<td>100 (33.3)</td>
<td></td>
</tr>
<tr>
<td>Age in years (mean, SD)</td>
<td>52.6 (7.6)</td>
<td>53.2 (8.0)</td>
<td>52.8 (7.2)</td>
<td>51.9 (7.5)</td>
<td>0.583</td>
</tr>
<tr>
<td>Age in years (median, IQR)</td>
<td>52 (13.0)</td>
<td>54 (14.0)</td>
<td>52 (11.0)</td>
<td>50.5 (11.3)</td>
<td></td>
</tr>
<tr>
<td>Marital status</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.001</td>
</tr>
<tr>
<td>Married with children</td>
<td>246 (82.0)</td>
<td>72 (72.0)</td>
<td>86 (86.0)</td>
<td>88 (88.0)</td>
<td></td>
</tr>
<tr>
<td>Married without children</td>
<td>14 (4.7)</td>
<td>8 (8.0)</td>
<td>2 (2.0)</td>
<td>4 (4.0)</td>
<td></td>
</tr>
<tr>
<td>Single</td>
<td>23 (7.7)</td>
<td>16 (16.0)</td>
<td>5 (5.0)</td>
<td>2 (2.0)</td>
<td></td>
</tr>
<tr>
<td>Divorced or separated</td>
<td>11 (3.7)</td>
<td>4 (4.0)</td>
<td>3 (3.0)</td>
<td>4 (4.0)</td>
<td></td>
</tr>
<tr>
<td>Widowed</td>
<td>6 (2.0)</td>
<td>–</td>
<td>4 (4.0)</td>
<td>2 (2.0)</td>
<td></td>
</tr>
<tr>
<td>Highest education completed</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Elementary school and below</td>
<td>48 (16.0)</td>
<td>8 (8.0)</td>
<td>23 (23.0)</td>
<td>17 (17.0)</td>
<td></td>
</tr>
<tr>
<td>High school sophomore or technical school</td>
<td>178 (59.3)</td>
<td>47 (47.0)</td>
<td>66 (66.0)</td>
<td>65 (65.0)</td>
<td></td>
</tr>
<tr>
<td>High school senior</td>
<td>46 (15.3)</td>
<td>29 (29.0)</td>
<td>10 (10.0)</td>
<td>7 (7.0)</td>
<td></td>
</tr>
<tr>
<td>College and above</td>
<td>28 (9.3)</td>
<td>16 (16.0)</td>
<td>1 (1.0)</td>
<td>11 (11.0)</td>
<td></td>
</tr>
<tr>
<td>Housing type</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.007</td>
</tr>
<tr>
<td>Small public housing</td>
<td>22 (7.3)</td>
<td>1 (1.0)</td>
<td>10 (10.0)</td>
<td>11 (11.0)</td>
<td></td>
</tr>
<tr>
<td>Medium public housing</td>
<td>188 (62.7)</td>
<td>62 (62.0)</td>
<td>68 (68.0)</td>
<td>58 (58.0)</td>
<td></td>
</tr>
<tr>
<td>Large public housing</td>
<td>82 (27.3)</td>
<td>33 (33.0)</td>
<td>19 (19.0)</td>
<td>30 (30.0)</td>
<td></td>
</tr>
<tr>
<td>Private apartment or landed property</td>
<td>8 (2.7)</td>
<td>4 (4.0)</td>
<td>3 (3.0)</td>
<td>1 (1.0)</td>
<td></td>
</tr>
<tr>
<td>Monthly household income</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.031</td>
</tr>
<tr>
<td>Below S$3,000</td>
<td>88 (29.3)</td>
<td>25 (25.0)</td>
<td>36 (36.0)</td>
<td>27 (27.0)</td>
<td></td>
</tr>
<tr>
<td>S$3,000–S$6,999</td>
<td>163 (54.3)</td>
<td>51 (51.0)</td>
<td>49 (49.0)</td>
<td>63 (63.0)</td>
<td></td>
</tr>
<tr>
<td>S$7,000 and above</td>
<td>49 (16.3)</td>
<td>24 (24.0)</td>
<td>15 (15.0)</td>
<td>10 (10.0)</td>
<td></td>
</tr>
<tr>
<td>Occupation</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Self-employed/PME</td>
<td>50 (16.7)</td>
<td>28 (28.0)</td>
<td>9 (9.0)</td>
<td>13 (13.0)</td>
<td></td>
</tr>
<tr>
<td>Other white-collar</td>
<td>73 (24.3)</td>
<td>32 (32.0)</td>
<td>15 (15.0)</td>
<td>26 (26.0)</td>
<td></td>
</tr>
<tr>
<td>Blue-collar</td>
<td>40 (13.3)</td>
<td>5 (5.0)</td>
<td>16 (16.0)</td>
<td>19 (19.0)</td>
<td></td>
</tr>
<tr>
<td>Housewife</td>
<td>121 (40.3)</td>
<td>28 (28.0)</td>
<td>53 (53.0)</td>
<td>40 (40.0)</td>
<td></td>
</tr>
<tr>
<td>Others, including retiree</td>
<td>16 (5.3)</td>
<td>7 (7.0)</td>
<td>7 (7.0)</td>
<td>2 (2.0)</td>
<td></td>
</tr>
<tr>
<td>Breast cancer related</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>0.092</td>
</tr>
<tr>
<td>Had a first-degree relative (biological mother, sister, or daughter) diagnosed with breast cancer before</td>
<td>13 (4.3)</td>
<td>7 (7.0)</td>
<td>1 (1.0)</td>
<td>5 (5.0)</td>
<td></td>
</tr>
<tr>
<td>Attended mammogram screening before</td>
<td>185 (61.7)</td>
<td>77 (77.0)</td>
<td>50 (50.0)</td>
<td>58 (58.0)</td>
<td></td>
</tr>
<tr>
<td>Time since last mammogram screening</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>&lt;1 year</td>
<td>47 (15.7)</td>
<td>22 (22.0)</td>
<td>9 (9.0)</td>
<td>16 (16.0)</td>
<td></td>
</tr>
<tr>
<td>1–2 years</td>
<td>57 (19.0)</td>
<td>24 (24.0)</td>
<td>13 (13.0)</td>
<td>20 (20.0)</td>
<td></td>
</tr>
<tr>
<td>2–3 years</td>
<td>29 (9.7)</td>
<td>8 (8.0)</td>
<td>9 (9.0)</td>
<td>12 (12.0)</td>
<td></td>
</tr>
<tr>
<td>3–5 years</td>
<td>13 (4.3)</td>
<td>4 (4.0)</td>
<td>6 (6.0)</td>
<td>3 (3.0)</td>
<td></td>
</tr>
<tr>
<td>&gt;5 years ago</td>
<td>39 (13.0)</td>
<td>19 (19.0)</td>
<td>13 (13.0)</td>
<td>7 (7.0)</td>
<td></td>
</tr>
</tbody>
</table>

**Note:** One-way ANOVA was used to test for significant differences in age between groups while chi-squared test was used to test for significant differences in categorical variables between groups.

**Abbreviations:** IQR, interquartile range; PME, professional, managerial, or executive; S$, Singapore dollars.
Table 3 Responder preferences (measured in PWUs) for attribute levels of SNP gene testing and interactions between respondent characteristics and attribute levels

<table>
<thead>
<tr>
<th>Attribute level*</th>
<th>Mean PWU (95% CI)</th>
<th>P-value</th>
<th>SD of mean PWU (95% CI)*</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Sample type</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Buccal swab</td>
<td>0.122 (0.002, 0.242)</td>
<td>0.046</td>
<td>0.290 (0.099, 0.481)*</td>
<td></td>
</tr>
<tr>
<td>Dried blood spot*</td>
<td>−0.122 (−0.242, −0.002)</td>
<td>0.046</td>
<td>−</td>
<td></td>
</tr>
<tr>
<td><strong>Person conducting pretest discussion</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Specialist doctor</td>
<td>1.032 (0.831, 1.233)</td>
<td>&lt;0.001</td>
<td>0.332 (0.099, 0.565)*</td>
<td></td>
</tr>
<tr>
<td>Nurse educator*</td>
<td>−0.443 (−0.620, −0.265)</td>
<td>&lt;0.001</td>
<td>−</td>
<td></td>
</tr>
<tr>
<td>Non-specialist doctor</td>
<td>−0.590 (−0.774, −0.406)</td>
<td>&lt;0.001</td>
<td>−0.336 (−0.553, −0.118)*</td>
<td></td>
</tr>
<tr>
<td><strong>Test location</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Private family clinic</td>
<td>0.158 (−0.019, 0.335)</td>
<td>0.080</td>
<td>0.273 (−0.019, 0.565)</td>
<td></td>
</tr>
<tr>
<td>Hospital*</td>
<td>−0.025 (−0.203, 0.152)</td>
<td>0.780</td>
<td>−</td>
<td></td>
</tr>
<tr>
<td>Public primary-care clinic</td>
<td>−0.133 (−0.304, 0.039)</td>
<td>0.130</td>
<td>−0.036 (−0.338, 0.266)</td>
<td></td>
</tr>
<tr>
<td><strong>Out-of-pocket cost, per S$100</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Choose left</td>
<td>−4.140 (−4.621, −3.660)</td>
<td>&lt;0.001</td>
<td>NA because coded as fixed</td>
<td></td>
</tr>
<tr>
<td>Choose none</td>
<td>−5.437 (−6.113, −4.762)</td>
<td>&lt;0.001</td>
<td>−0.318 (−0.545, −0.091)*</td>
<td></td>
</tr>
<tr>
<td><strong>Alternative-specific constants</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Significant interactions in the main effects + interactions model*</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Age 60–69 × choose left</td>
<td>−0.384 (−0.623, −0.145)</td>
<td>0.002</td>
<td>−0.303 (−0.596, −0.010)*</td>
<td></td>
</tr>
<tr>
<td>Income $33,000–$69,999 × choose left</td>
<td>−0.251 (−0.453, −0.050)</td>
<td>0.015</td>
<td>0.573 (0.344, 0.803)*</td>
<td></td>
</tr>
<tr>
<td>Income &gt;$70,000 × cost per S$100</td>
<td>0.586 (0.178, 0.995)</td>
<td>0.005</td>
<td>1.382 (0.987, 1.777)*</td>
<td></td>
</tr>
<tr>
<td>High school senior education × cost per S$100</td>
<td>0.917 (0.469, 1.366)</td>
<td>&lt;0.001</td>
<td>−0.181 (−0.799, 0.438)</td>
<td></td>
</tr>
<tr>
<td>Blue-collar occupation × cost per S$100</td>
<td>−0.694 (−1.384, −0.004)</td>
<td>&lt;0.001</td>
<td>0.846 (0.457, 1.236)*</td>
<td></td>
</tr>
</tbody>
</table>

Notes: Other variables in the model include (not significant): Indian ethnicity × cost per S$100, age 50–59 × choose left, age 50–59 × cost per S$100, age 60–69 × cost per S$100, white-collar occupation × cost per S$100, and other occupation × cost per S$100. Reference levels for respondent characteristics are: age 40–49 years, household income <$33,000, elementary school education, self-employed/professional, managerial, or executive, and Chinese ethnicity. Negative PWU represents disutility. Grand mean has an expected utility of zero. *All parameter estimates were derived from the mixed logit model with main effects plus interactions estimated in the preference space, as it was a better fit for the data as compared to main effects model (likelihood ratio test, P<0.001). †The sign of the estimated standard deviations is irrelevant and should be interpreted as being positive. The likelihood ratio test for the joint significance of all standard deviations has a P-value of <0.001, implying that the null hypothesis that all standard deviations are equal to zero is rejected, that is, there is significant preference heterogeneity. ‡Signifies the reference level of each attribute. Parameter estimates of reference levels were obtained using the lincom command in Stata. §P<0.05. ‡P<0.01. $P<0.001.

Abbreviations: NA, not applicable; PWU, part-worth utility; SNP, single-nucleotide polymorphism; S$, Singapore dollars.

only for out-of-pocket cost and ASC “Choose left.” Those with household income >$7,000 and high school senior education have overall lesser disutility toward out-of-pocket cost (P=0.005 and <0.001, respectively), while blue-collar occupation attached greater disutility to cost (P=0.049). This implies that higher income women and those holding professional, managerial, or executive (PME, reference level) jobs are more willing to pay out-of-pocket than lower income women and blue-collar workers. Women with household income between $3,000 and $6,999 and women aged between 60 and 69 years attached lesser utility toward the left-sided alternative (P=0.015 and 0.002, respectively). Except for private family clinic, public primary-care clinic, and the interaction term between high school senior education and cost, we observed that the estimated SD coefficients for all other attribute levels are significant, indicating the presence of preference heterogeneity (Table 3).

Probability density and cumulative density functions of individual mWTP are presented in Figure 2. Women are willing to pay to transit to buccal swab, specialist doctor, and to a lesser degree private family clinic, and not willing to pay for non-specialist doctor and public primary-care clinic, where the mWTP values are less than zero. Respondents are willing to pay the most for the transition from nurse educator to specialist doctor at $28.39 (95% CI $25.35, $31.42) and least willing to pay for the transition from nurse educator to non-specialist doctor at −$16.53 (95% CI −$19.18, −$13.88) (Table 4).

Subgroup analyses

Preferences among ethnic groups are represented by mWTP distributions in Figure 3 and summarized in Table 5. In general, they seem to be rather similar. With reference to the probability density functions, the Indians have narrower density distribution for buccal swab, concentrated around mWTP of $2.66. Together with the Malays, their densities are slightly shifted toward higher mWTP than the Chinese. For specialist doctor, the Malays have lower mWTP, depicted by the higher density than Chinese and Indians around mWTP of $15–$25. For non-specialist doctor, the Chinese appear
to have a much longer right tail toward higher mWTP, although it is still in the negative range. They also appear to be more willing to pay for private family clinic, with higher density around mWTP of S$13–S$25. Cumulative density functions for all attributes seem to be comparable except for private family clinic, where the median mWTP of Chinese is S$8.31 as compared to S$6.99 in Malays and S$7.63 in Indians (Table 5). Chinese women have greater preference heterogeneity for the attributes as compared to the other two ethnicities, judging by the longer tails at both ends of the probability density functions.

**Predicted uptake rates in various scenarios**

Predicted uptake rates for various scenarios are shown in Table 6. A realistic base case scenario consisted of less invasive testing (buccal swab) involving a pretest discussion with a specialist doctor at the hospital which costs S$175.

![Figure 2](Continued)
Other scenarios include a feasible alternative that substitutes specialist doctor with nurse educator at a lower fee, cheaper alternative at S$50, pricier alternative at S$300, and the alternative consisting of the most preferred attribute levels at the population level (n=300). As compared to the base case, predicted uptake increase from 3.7% to 60.5% in feasible alternative and increase to 87.0% in cheaper alternative and 88.9% in most preferred alternative. Uptake decrease to

Table 4 Overall respondent mWTP for attribute levels of SNP gene testing (n=300)

<table>
<thead>
<tr>
<th>Attribute level</th>
<th>Mean mWTP, a S$ (95% CI)</th>
<th>P-value</th>
<th>SD of mean mWTP, b S$ (95% CI)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample type</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Buccal swab</td>
<td>2.65 (0.79, 4.51)</td>
<td>0.005</td>
<td>1.48 (-0.69, 3.65)</td>
</tr>
<tr>
<td>Dried blood spot</td>
<td>–</td>
<td>–</td>
<td>–</td>
</tr>
<tr>
<td>Person conducting pre-test discussion</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Specialist doctor</td>
<td>28.39 (25.35, 31.42)</td>
<td>&lt;0.001</td>
<td>10.87 (7.93, 13.78)</td>
</tr>
<tr>
<td>Nurse educator</td>
<td>–</td>
<td>–</td>
<td>–</td>
</tr>
<tr>
<td>Non-specialist doctor</td>
<td>-16.53 (-19.18, -13.88)</td>
<td>&lt;0.001</td>
<td>-2.91 (-5.05, -0.77)</td>
</tr>
<tr>
<td>Test location</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Private family clinic</td>
<td>7.70 (4.22, 11.17)</td>
<td>&lt;0.001</td>
<td>15.39 (11.94, 18.84)</td>
</tr>
<tr>
<td>Hospital</td>
<td>–</td>
<td>–</td>
<td>–</td>
</tr>
<tr>
<td>Public primary-care clinic</td>
<td>-5.86 (-8.94, -2.79)</td>
<td>&lt;0.001</td>
<td>-1.43 (-5.43, 2.57)</td>
</tr>
</tbody>
</table>

Notes: a mWTP figures were obtained from the same mixed logit model in Table 3 but estimated in WTP space (main effects plus interactions). b The sign of the estimated standard deviations is irrelevant and should be interpreted as being positive. The likelihood ratio test for the joint significance of all standard deviations has a P-value of <0.001, implying that the null hypothesis that all standard deviations are equal to zero is rejected, that is, there is significant preference heterogeneity. c Signifies the reference level of each attribute. d (P<0.01), e (P<0.001).

Abbreviations: mWTP, marginal willingness-to-pay; SNP, single-nucleotide polymorphism; S$, Singapore dollars.
almost 0% in pricier alternative, where cost is the highest at S$300.

**Dual-response none**

One hundred of 300 women (33.3%) answered “No” to “Will you really go for this option in real life?” at least once in 10 choice tasks; 10 in 300 (3.3%) women always opted out of the SNP gene test in real life; 333 out of 3,000 choice tasks (11.0%) are answered as “No.” Reasons include high cost (n=223), unaware of such tests (n=72), not keen to know the results (n=10), lack of time (n=10), afraid of taking blood (n=7), currently healthy therefore test is unnecessary (n=6), person conducting pretest discussion is not a specialist doctor (n=3), and others (n=2).

**Discussion**

This is the first study to quantify women’s preferences, mWTP, and predicted uptake for SNP gene testing in breast cancer. Our findings highlight the importance of incorporating culturally and linguistically appropriate services,
as advocated by the National Committee of Quality Assurance which focuses on improving the quality of healthcare. We found that person conducting the pretest discussion and cost are significantly associated with women’s choice of getting tested, while sample type (noninvasive buccal swab preferred) is only marginally significant ($P=0.046$). While preference for pretest discussion was not previously investigated, a recent DCE found that patients with chronic lymphocytic leukemia prefer to receive genomic test results from a hospital consultant doctor or specialist nurse rather than general practitioner or junior doctor. Similar to our findings, cost influences the attendance of cancer screening as well as pharmacogenetic test that predicts risk of life-threatening adverse drug reactions. Respondents also preferred noninvasive procedures in genome-based colorectal cancer screening scenarios.

Although we observed a bias toward the left-sided alternative, there is little difference between preference weights of attribute levels including and excluding the ASC “Choose left.” Nonetheless, we considered the possibility that left-sided alternatives are consistently superior throughout the questionnaire. This is not found to be true when we checked through the design. Respondents could have been influenced by the first (fixed) task where option A was superior. Alternatively, they might have applied simplifying heuristics when undertaking this survey. However, we tried to minimize the application of heuristics by getting interviewers to emphasize to the respondents that either option may appeal more to them. We also did not observe any serial non-traders who always chose the left-sided option in every choice task.

Through the interactions analyses, we found that women with higher household income and education level were more willing to pay higher prices for SNP gene testing, while blue-collar workers were less willing than the self-employed or PME. This is in line with several studies demonstrating association between socioeconomic factors, for example, income and education on screening test uptake. We observed significant preference heterogeneity for most attribute levels, suggesting the importance of tailoring SNP gene testing packages according to individual preferences, to maximize utility. Both the mean PWU and SD coefficients for buccal swab were significant at 5% level. By calculating the cumulative

![Figure 3](https://www.dovepress.com/a.png)

**Figure 3** Probability density and cumulative density functions of mWTP across three ethnic groups. **Abbreviation:** mWTP, marginal willingness-to-pay.
density function under the standardized normal curve, we
find that 66% of the distribution is above zero, while 34% is
below zero. This implies that buccal swab is associated with
positive utility among about two-thirds of the respondents.
The policy implication could be to provide buccal swab to
those who self-select this option. Although preferences in
terms of mWTP across ethnic groups appear to be compa-
rable, there are subtle differences in mWTP for private family
clinic and extent of preference heterogeneity. Nonetheless,
cost-effectiveness of a one-size-fits-all strategy versus a
personalized strategy should be empirically studied.

Women were willing to pay more to discuss with a spe-
cialist doctor, but not with a non-specialist doctor. However,
this may not be the best use of specialist doctor’ time. Since
we found that nurse educator was preferred after specialist
doctor, they can be considered as alternatives to specialist
doctors. Research has shown that nurses trained in genetic
counseling are well-positioned in providing such services
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counseling are well-positioned in providing such services
professionals as genetic counselors. We believe that women will be receptive toward genetic counselors if their proficiency and cost-effectiveness are demonstrated clearly.

As for predicted uptake, the feasible alternative in Table 6 appeared to be relatively well-received, achieving rates similar to that of a pharmacogenetic test to prevent severe adverse reactions and genome-based colorectal cancer screening. The cheaper and most preferred alternative may also become feasible in the future, as the cost of multi-panel gene testing becomes more affordable. The two most common reasons for opting out of the test in real life were high cost (n=233) and lack of awareness of such tests (n=72). This shows that the end-user’s sensitivity toward price and knowledge of the health technology can influence its subsequent uptake. Nonetheless, only 3.3% of respondents always opted out, as compared to 21.4% who opted out of diabetes lifestyle management programs and 29.0% who opted out of medically supervised benzodiazepine discontinuation programs.

There are some limitations in this study. First, we assumed that both alternatives are equally highly accurate. Our study conclusions may change, depending on the clinical validity and utility of SNP gene testing. Second, the actual uptake of breast cancer screening in Singapore is rather low. Only about 40% of women aged 50–59 years and 38% of women aged 60–69 years surveyed in 2010 who participated in screening in the past 2 years. Similarly, only about 35% of our study sample attended mammogram screening within the past 2 years. Actual uptake of SNP gene testing may be even lower, as it only predicts the risk and not the actual occurrence of breast cancer. Nonetheless, SNP gene testing holds the promise of reducing screening harms for those at lower risk and detecting cancers earlier for those at higher risk. A large-scale prospective, randomized controlled trial is underway to determine whether screening based on personalized risk is as safe, less morbid, and more preferred by women than age-based screening. Other limitations include the uncertainty surrounding the use of heuristics or simplifying tactics to make decisions. Due to the nature of our sampling and recruitment, our study population tend to comprise housewives and women who were not working. They might have received lower education and have lower health literacy as compared to the general women population between 40 and 69 years. In addition, we could not reach out to women staying in condominiums or private apartments who earn higher incomes and may have higher health literacy. Last but not least, almost 50% of the eligible households did not participate in the study. These women may be less concerned with health-related issues and less willing to undergoing SNP gene testing in real life. Thus, our study findings may not be generalizable to them.

Conclusion
Sample type, person conducting pretest discussion, and out-of-pocket cost are significantly associated with preference for SNP gene testing. Women with higher income and education levels were more willing to pay higher prices for the test. SNP gene testing using buccal swab involving a pretest discussion with a trained nurse educator at the hospital is more preferred and may be more cost-effective than the base case scenario involving specialist doctor at a higher cost. It could be personalized according to individual preferences to encourage uptake. Future research should focus on establishing the clinical validity and utility of SNP gene testing, the long-term outcomes of risk-based screening and prevention strategies, in terms of patients’ experiences and survival as well as the cost-effectiveness of tailored breast cancer screening using SNP gene testing.

Data availability
The data sets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

Acknowledgments
We thank A/Prof Alex Cook, PhD (Saw Swee Hock School of Public Health) for providing editorial assistance as well as two anonymous reviewers for suggesting substantial improvements. This study was supported by the Singapore Ministry of Health Health Services Research Competitive Research Grant, administered by National Medical Research Council (Grant number: HSRG/13MAY006). The study sponsor had no role in the study design, collection, analysis, and interpretation of data; in the writing of the manuscript; and in the decision to submit the manuscript for publication.

Disclosure
The authors report no conflicts of interest in this work.

References


Supplementary materials
Additional information on study methods
Multi-stage cluster sampling
A market research company was commissioned to recruit respondents and implement the survey. Geographical clusters were first selected from the company’s own residential sampling frame using systematic sampling. Due to access difficulties, condominium and private apartment dwellers (13.9% of the population) were excluded. Households in each Primary Sampling Unit were then selected door-to-door. Only one respondent per household could participate in the study.

Discrete choice experiment (DCE) survey development
To fully capture women’s views and beliefs toward single-nucleotide polymorphism gene testing to guide personalized screening strategies, focus group discussions involving screening-age Singaporean women were conducted. Data saturation was reached after four sessions. It was found that accuracy, invasiveness, cost, and side effects influence women’s uptake of the test.

A pilot DCE survey was then conducted in 20 respondents, recruited in the same manner. Some respondents considered cousins as first-degree relatives in the question on family history of breast cancer. As such, we modified the question to read “has your biological mother, sister or daughter been diagnosed with breast cancer before?” Three out of 20 respondents did not understand the term “nurse educator.” Hence, we included an explanation of the role of a nurse educator in the two-page write-up.

Statistical analyses
Predicted uptake can be measured using the following formula:

\[
P(\text{accepting alternative } i) = \frac{1}{1 + e^{-V_i}}
\]

where \(V_i\) is the expected utility of alternative \(i\), or the sum of part-worth utilities for the attributes corresponding to the specific levels. An alternative with expected utility \(V\) equals to zero is assumed to have a 50% probability of acceptance. Due to effects coding, all parameters are estimated relative to the grand mean, which has an expected utility of zero.

Table S1 Comparison between PWUs of attribute levels of SNP gene testing including and excluding alternative-specific constant “choose left” (main effects only)

<table>
<thead>
<tr>
<th>Attribute level</th>
<th>Mean PWU (95% CI)</th>
<th>Excluding “Choose left”</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Including “Choose left”</td>
<td></td>
</tr>
<tr>
<td>Buccal swab</td>
<td>0.087 (−0.017, 0.190)</td>
<td>0.070 (−0.014, 0.154)</td>
</tr>
<tr>
<td>Dried blood spot(^a)</td>
<td>−0.087 (−0.190, 0.017)</td>
<td>−0.070 (−0.154, 0.014)</td>
</tr>
<tr>
<td>Specialist doctor</td>
<td>0.776 (0.617, 0.936)(^c)</td>
<td>0.522 (0.398, 0.646)(^d)</td>
</tr>
<tr>
<td>Nurse educator(^a)</td>
<td>−0.327 (−0.474, −0.180)(^s)</td>
<td>−0.251 (−0.377, −0.125)(^s)</td>
</tr>
<tr>
<td>Non-specialist doctor</td>
<td>−0.449 (−0.602, −0.297)(^s)</td>
<td>−0.271 (−0.387, −0.156)(^s)</td>
</tr>
<tr>
<td>Private family clinic</td>
<td>0.091 (−0.056, 0.238)</td>
<td>0.084 (−0.037, 0.205)</td>
</tr>
<tr>
<td>Hospital(^a)</td>
<td>0.029 (−0.124, 0.183)</td>
<td>−0.074 (−0.205, 0.057)</td>
</tr>
<tr>
<td>Public primary-care clinic</td>
<td>−0.120 (−0.271, 0.031)</td>
<td>−0.010 (−0.133, 0.112)</td>
</tr>
<tr>
<td>Out-of-pocket cost, per S$100</td>
<td>−2.848 (−3.082, −2.613)(^s)</td>
<td>−1.840 (−1.972, −1.709)(^s)</td>
</tr>
<tr>
<td>Alternative-specific constant(s)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Choose left</td>
<td>2.092 (1.887, 2.296)(^d)</td>
<td>NA</td>
</tr>
<tr>
<td>Choose none</td>
<td>−3.801 (−4.163, −3.439)(^s)</td>
<td>−1.770 (−1.917, −1.623)(^s)</td>
</tr>
</tbody>
</table>

Notes: Negative PWU represents disutility. Grand mean has an expected utility of zero. \(^s\)Signifies the reference level of each attribute. Parameter estimates were obtained using the lincom command in Stata. \(^P<0.001\).

Abbreviations: NA, not applicable; PWU, part-worth utility; SNP, single-nucleotide polymorphism.
References